COINCIDENCE OF CROSSING OVER IN DROSOPHILA MELANOGASTER (AMPELOPHILA)¹

ALEXANDER WEINSTEIN

Station for Experimental Evolution, Cold Spring Harbor, New York.

[Received May 24, 1917]

TABLE OF CONTENTS

	PAGE
INTRODUCTION	. 135
Coincidence of widely separated regions	. 138
Coincidence in the second chromosome	. 146
Coincidence in the third chromosome	. 148
Coincidence in other forms	. 148
The mechanism of crossing over	. 148
Triple crossing over	. 149
Maximum and minimum coincidence	. 153
The distance between the breaking points in double crossing over	. 154
Mutations observed	. 156
Summary	. 158
LITERATURE CITED	. 158
TABLES	. 160

The present investigation was undertaken in the hope of throwing some light on the behavior of the chromosomes during crossing over. The problem was considered from the point of view of genetics by studying the effect of crossing over in one region of the chromosome on crossing over in another region, and by considering how this effect varies with variation of the distance between the regions involved. The phenomena have a bearing on the method of twisting of the chromosomes and the mechanism of crossing over.

According to the theory of linear arrangement and the chiasmatype, Mendelian genes are disposed in linear series in the chromosomes, and the separation of linked genes (crossing over) is due to breaks in homologous chromosomes (which have come together during synapsis) and recombinations between the resulting pieces (JANSSENS 1909, MOR-GAN 1910, STURTEVANT 1913, 1915, MULLER 1916). The percentage of cases in which two linked genes separate (amount of crossing over between them) is necessarily proportional, other things being equal, to

¹ Contribution from Zoölogical Laboratory of Columbia University. Genetics 3: 135 Mr 1918

ALEXANDER WEINSTEIN

the distance between the genes. Hence the distance on a chromosome map is measured by this percentage of crossing over. But this percentage is also affected by other conditions than actual morphological distance; for example, the age of the female, definite Mendelian factors, etc. (STURTEVANT 1913, 1915, BRIDGES 1915, MULLER 1916). Moreover, different regions may be, and in some cases are known to be, affected to different extents. The distances on chromosome maps are, therefore, not necessarily to be interpreted as actual morphological distances, for it is possible that the amount of crossing over in the morphologically shorter of two distances is greater than in the morphologically longer. Therefore, when the term *distance* is used in this paper, it should be interpreted as *percentage of crossing over*.

Breaks may occur simultaneously at more than one point of a chromosome. For example, in the disjunction of two chromosomes of which one contains the factors $M \ N \ P \ Q$ and the other the allelomorphic factors $m \ n \ p \ q$ (see figure 1), when a break occurs between M and N a break may also occur between P and Q. If the pieces recombine, the resulting chromosomes will contain respectively the factors $m \ N \ P \ q$ and $M \ n \ p \ Q$.



It was evident, however, from the data on which the theories above referred to were based, that when a break occurs in one region of a chromosome, neighboring regions are much less likely to undergo a break

than would otherwise be the case. If, in the above example, the regions M N and P Q are not far apart, a break between M and N will tend to prevent the occurrence of a break between P and Q (STURTEVANT 1913, 1915, MORGAN, STURTEVANT, MULLER and BRIDGES 1915, pp. 63-64, MULLER 1916).

This phenomenon, which has been termed *interference*, is well illustrated by an unpublished cross carried out by BRIDGES. The cross involved four sex-linked factors in Drosophila: vermilion (eye color), sable (body color), garnet (eye color) and forked (bristles). The relative positions of these factors are shown in the diagram (figure 2).



The amount of crossing over in the vermilion sable region is about 10 percent; in the garnet forked region about 12 percent. That is, if we take all the offspring from this cross, about 10 percent will be crossovers between vermilion and sable, and about 12 percent will be crossovers between garnet and forked. But if instead of taking all the flies we take only those which have already resulted from a crossing over between vermilion and sable, we find that in this sample the proportion that are also crossovers between garnet and forked is not 12 percent but only 1.2 percent. That is, these crossovers are only one-tenth as numerous as in a random sample. Or, if we count only the crossovers between vermilion and sable is not 10 percent, but only 1.0 percent—again only one-tenth as many as in a random sample.

The actual data are as follows:1

				TABLE I	•			
	T	2	3	12	<u>1 3</u>	23	123	Total
2651	313	47	380	0	3	0	0	3394

Since the total amount of crossing over in the vermilion sable region is 0.093 and in the garnet forked region 0.113, the amount of crossing

¹ In this and other tables the non-crossover class is denoted by a line, without numerals. The crossover classes are denoted by a line and numerals referring to the regions in which the crossing over took place, the regions being numbered from left to right. In the present case, the vermilion sable region is **1**, the sable garnet region is **2**, the garnet forked region is **3**. \underline{I} — indicates a crossing over in the vermilion sable region, \underline{I} — $\underline{3}$ indicates a double crossing over involving the vermilion sable and garnet forked regions, etc. In the text, crossover classes may be indicated by the numbers of the regions in which crossing over in both the first and third regions.

GENETICS 3: Mr 1918

over involving both regions simultaneously (if they were independent of each other) would be $0.093 \times 0.113 = 0.01$. The observed proportion of crossing over involving both regions is only 0.001. The ratio of the actual amount to the amount expected if the regions did not affect each other is $\frac{0.001}{0.01} = 0.10$. This ratio is termed the *coincidence of crossing over* of the two regions (BRIDGES 1915, MULLER 1916).² That is, the coincidence of crossing over of two regions is the ratio of the actual amount of crossing over involving both regions to the amount that would occur if crossing over in one region did not affect crossing over in the other.³

It is evident that in a double crossing over involving the regions M Nand P Q (see figure 3), the distance between the two points at which

FIGURE 3

the breaks occur may be just greater than N P or just less than M Q; or it may have any intermediate value. It can be shown that the average distance between the breaking points is the mean of N P and M Q, or the distance from the mid-point of M N to the mid-point of P Q. This was pointed out by MULLER (1916). The proof of this and a discussion of the assumptions on which it rests will be given later in this paper; I wish here merely to call attention to the formula.

COINCIDENCE OF WIDELY SEPARATED REGIONS

In the vermilion garnet sable forked cross above cited, the intermediate distance (sable garnet) is very short—only about two units. The work on Drosophila has shown that in general the coincidence of two regions increases as the distance between them increases (STURTEVANT 1915, MULLER 1916). The present investigation was undertaken to determine the value of coincidence when the intermediate distance becomes very long. Two crosses were made involving factors in the sex chromosome. The relative positions of these factors are shown in figure 4.



² This phenomenon was originally referred to in terms of interference. The index of interference used was the reciprocal of coincidence (STURTEVANT 1913, 1915, BRIDGES 1915, MULLER 1916).

⁸ MULLER has preferred to state this ratio as a percentage; but for greater ease in calculation it is better expressed as a decimal fraction. The latter usage will be adopted in this paper.

One cross involved the factors eosin (eye color), ruby (eye color), sable (body color) and forked (bristles); and the other cross involved eosin, ruby, forked and cleft (venation). Practically the entire length of the sex chromosome (so far as known) is involved in the latter cross, since eosin is only one unit from the extreme left-hand end and cleft is (with the possible exception of lethal sc) the extreme right-hand factor. The farthest right-hand factor hitherto worked with in determinations of coincidence has been bar, which is only half a unit to the right of forked.

In choosing the factors to be worked with, it was necessary (for reasons that will be explained later) to make the regions whose coincidence was to be determined short enough to prevent the occurrence of double crossing over within them. This made the percentage of crossing over within each region small and necessitated making large counts to get significant results.

The results indicate that the coincidence of the eosin ruby region and the sable forked region (a mean intermediate distance of about 46) is about 1.00 (possibly a little more), and that the coincidence of the eosin ruby region with the forked cleft region (a mean intermediate distance of about 57) is only 0.68. That is, when the intermediate distance increases beyond a certain value, coincidence instead of rising or remaining at the value it has reached, falls again. This means that crossing over at the extreme left-hand end of the chromosome interferes with crossing over at the right-hand end, but has practically no effect on crossing over in the region just to the left of the right-hand end.

In each cross, females heterozygous for the particular combination of factors under consideration were bred singly. Since the factors involved are sex-linked, the proportion of the different classes of sons is a direct index of the frequency of crossing over in the different regions of the chromosome. The daughters were counted in order to keep track of the sex ratios, but were not classified, because it is often impossible to distinguish eosin from ruby in the female, and because it is impossible to backcross by cleft males, which are sterile. The mothers, therefore, instead of being backcrossed, were in most cases mated to bar males. Bar is a dominant sex-linked character; hence it was possible to distinguish non-disjunction in both male and female offspring (BRIDGES 1916). Cultures giving non-disjunction were excluded from the totals in both crosses, because of the possibility that non-disjunction might affect coincidence.

GENETICS 3: Mr 1918

ALEXANDER WEINSTEIN

Each cross was made so that the factors involved entered in several different combinations in different matings. In this way any particular crossover class is represented by several different combinations of characters, so that if the viability of any particular combination of characters is poor, it may be counterbalanced by the good viability of another combination and *vice versa* (MORGAN and BRIDGES 1916). Moreover, cultures giving less than 40 males were excluded because a small number of offspring is often due to poor culture conditions, which may entail differential inviability.

The data for the eosin ruby sable forked cross are as follows:

					Table	2			
φ φ	$\delta \delta$								Total රී රී
		I	2	3	12	<u>13</u>	23	123	
15298	6363	579	3638	1208	128	115	350	6	12387

The total proportionate amount of crossing over in the eosin ruby region is here 0.0668. The total amount of crossing over in the sable forked region is 0.1355. If crossing over in one region does not affect crossing over in the other, the amount of crossing over involving both regions simultaneously would be $0.0668 \times 0.1355 = 0.00905$. The actual amount of double crossing over involving both regions (the triple crossovers are excluded for reasons that will be explained later) is 0.00928. The coincidence is $\frac{0.00928}{0.00905} = 1.025$.⁴ The two regions are therefore independent. If anything, a crossing over between eosin and ruby makes crossing over between sable and forked slightly more likely

to occur. The counts of the eosin ruby forked cleft cross are given in the fol-

The counts of the eosin ruby forked cleft cross are given in the following table:

⁴ The actual calculation of coincidence can be made somewhat more simply. For if the total number of flies involved is n, the total number of crossovers in the two regions respectively a and b, and the total number of double crossovers involving both regions x, then

the coincidence
$$=$$
 $\frac{\frac{x}{n}}{\frac{a}{n} \times \frac{b}{n}} = \frac{xn}{ab}$

This makes it possible to calculate the coincidence of two regions without calculating their lengths.

The total amount of crossing over in the eosin ruby region is 0.0560; the amount of crossing over in the forked cleft region is 0.0875. The amount of double crossing over involving both regions (if they do not affect each other) should be $0.0560 \times 0.0875 = 0.004904$. The actual amount of double crossing over involving both regions is 0.003316. The 0.003316

coincidence is $\frac{0.003316}{0.004904} = 0.6761$. This indicates that crossing over in

the eosin ruby region hinders crossing over in the forked cleft region.

The value obtained for the coincidence in each case is, of course, subject to a probable error of random sampling. But the applicability of the ordinary formula is doubtful in cases where the class under consideration (here the I, 3 double crossover class), forms as small a proportion of the total as in the present case. PEARL (1917) has therefore suggested another method. DR. PEARL was kind enough to undertake a calculation of the significance of the present data, which was published in the paper referred to. This calculation, based on all the data in the eosin ruby sable forked cross, but only 9017 flies in the eosin ruby forked cleft cross, gives a probability of 0.0421; that is, the chances are about 23 to I against obtaining in the eosin ruby sable forked cross a coincidence as low as, or lower than, that obtained in the eosin ruby forked cleft cross.⁵

Since more data are now available, a recalculation on the basis of all the data has been made. This gives for the lower quartile of the 1,3 double crossover class in the eosin ruby sable forked cross, the value 65.3951, which is considerably higher than 47. But the probability to which this value corresponds cannot be determined from the table given by PEARL, since the latter is based on only part of the data involved.

On the Gaussian hypothesis, the observed deviation would correspond to odds of about 105 to I that the two values are different. But since the Gaussian method is supposed to be inapplicable to cases where p is very small and since statisticians are not in agreement on the validity of the other method (PEARSON 1917; compare also WRIGHT 1917), it seems best to suspend judgment on the statistical significance of the data.

There is, moreover, an objection to calculating the coincidence, as has just been done, on the basis of all the data lumped together. This is due to the fact that in any series of matings (such as the cosin ruby sable

⁵ In the data submitted to Dr. PEARL, I did not take into account the difference between the values of the cosin ruby distance obtained in the two crosses; hence the probability arrived at by Dr. PEARL is different from that given here. This fact, of course, in no way affects the validity of the method used.

forked cross) the coincidence might be the same in all the cultures; and yet if the coincidence were calculated from the total data instead of for each culture separately, the value derived might be different from that for each separate culture, if the amount of crossing over varied in the separate bottles.

This will be evident from a simple numerical example. Let us take two imaginary cultures of 100 flies each. Let us suppose that in one culture the number of crossovers in each of the two regions involved is 10 (or 0.10 of the total) and the number of double crossovers involving both regions is 1 (or 0.01 of the total). Let us suppose that in the second culture the number of crossovers in each of the regions is 20 (or 0.20 of the total) and the number involving both regions at once is 4 (or 0.04 of the total). The coincidence in the first case is $\frac{0.01}{0.01}$ = 1.00; in the second case it is $\frac{0.04}{0.04}$, which is also 1.00; but if we calculate

the coincidence on both cultures together it is $\frac{200}{\frac{30}{200}} = 1.11.$

We can state this in more general terms. In one culture let *m* be the total number of flies, *a* and *b* the number of crossovers involving respectively the two regions under consideration, and *x* the number of double crossovers involving both regions simultaneously. In another culture let the corresponding values be *n*, *p* and *q*, and *y*. Then the co-incidence in the first case (see footnote 4, p. 140) is $\frac{xm}{ab}$ and in the second case $\frac{yn}{pq}$. The average of these two values is $\frac{y(xm)}{pq} + \frac{yn}{pq} = \frac{xmpq + ynab}{2abpq}$. The coincidence calculated on the total data is $\frac{(x+y)(m+n)}{(a+p)(b+q)}$. It is evident that these two values are not necessarily equal.

The coincidence in each cross was therefore calculated by considering each culture separately. The values obtained are given in the tables at the end of the paper. (In a few cultures there was no crossing over in

one of the two regions involved. In such cases the value of the coincidence is $\frac{o}{o}$, which is indeterminate. These cultures are consequently omitted from the calculation.)

The average value of the coincidence of eosin ruby and sable forked calculated in this way is 1.0238. The average value of the coincidence of eosin ruby and forked cleft is 0.6064. If the values are weighted according to the number of individuals in each culture, the averages are practically the same; namely, 1.0081 and 0.6049 respectively. These values agree very well with the values obtained in the previous calculation.

In table 4 the values of coincidence are grouped by intervals of 1.00, and the accompanying curves (figures 5 and 6) show these distributions graphically.

Coincidence	0	I	2	3	4	5	6	7	8	9	10	Total broods
Number of broods, eosin ruby sable forked cross	91	41	12	6	2	4	I	0	0	0	0	157
Number of broods, eosin ruby forked cleft cross	141	13	7	4	3	2	I	I	0	0	I	173

TABLE 4

It is noticeable that there is a comparatively smaller number of cultures whose coincidence is zero and a greater number of cultures whose coincidence is between I and 2 in the eosin ruby sable forked cross than in the eosin ruby forked cleft cross. This is partly at least due to the fact that the expected number of double crossovers involving the particular regions under consideration would (quite apart from considerations of coincidence) be greater in the former cross than in the latter, because the distance between sable and forked is greater than that between forked and cleft, and also because the eosin ruby distance in the first cross is slightly longer than in the second. However, this is counterbalanced to some extent at least by the fact that when a double crossover does occur in the second cross it raises the coincidence relatively more than in the first cross. It is in accordance with this that the highest values for coincidence are found in cultures of the second cross, not of the first, although the average coincidence in the first is higher.

Bearing in mind the possibilities of error due to these facts, we may

GENETICS 3: Mr 1918

apply the χ^2 test (PEARSON 1911) to these two distributions. We obtain for χ^2 a value of 29.166. By referring to ELDERTON's tables (ELDERTON 1901, p. 159) we find that this value for n = 9 corresponds to a chance



FIGURE 5.—Distribution of the values of coincidence of eosin ruby and sable forked.



of 0.000299. Since this is equivalent to only one chance in 3344, the chance that the two distributions are the same is statistically negligible.

In some of the matings involving eosin, ruby, forked and cleft, a fifth factor, lozenge (eye), was also followed. Figure 7 indicates the location of this factor with respect to the others.



The intermediate distance between the ruby lozenge and the forked cleft regions is about the same as that between the eosin ruby and the sable forked regions. Hence we might expect the coincidence of ruby lozenge and forked cleft to be about the same as that of eosin ruby and sable forked, and greater than that of eosin ruby and forked cleft. The actual counts in the present cross are as follows:

								TAB	LE 5							
ęφ	ත් ත්															Total or 50
		I	2	3	4	12	13	14	23	24	34	123	124	<u>134</u>	234	
3409	1217	96	321	599	164	3	31	4	53	34	44	I	I	3	I	2572

While these data are not sufficiently extensive in themselves to be significant, it may be instructive to compare them with the data already considered. The coincidence of ruby lozenge and forked cleft is here 0.866; that of eosin ruby and forked cleft is 0.295. Both values are lower than those obtained for the similar distances in the other crosses; but the disproportion between them is in the same direction and is even more marked.

A cross was also made with	the factors eosin, rul	oy, forked and	d fused
(wing). Since fused is betw	een forked and cleft	(figure 8),	the co-
eosin ruby	sable	forked fused	cl e ft
		1 1	1
I 7	43	56.5 59.5	65
	FIGURE 8		

incidence of eosin ruby and forked fused might be expected to be intermediate between that of eosin ruby and sable forked and that of eosin ruby and forked cleft. The data obtained were as follows:

				Т	'able 6				
ŞΥ	රිරි								Total ර්ර්
		1	2	3	I 2	<u>13</u>	23	123	•
8220	3889	290	3027	103	130	7	59	I	7506

Here the proportion of crossing over between eosin and ruby is 0.5702; between forked and fused the proportion is 0.0226. The coincidence of these two regions is 0.7221. This agrees with the expectation.

The cultures included above, as in the other crosses, were only those which contained at least 40 males, but since the number was too small to be significant another table was made in which all the other available cultures were also added in. The resulting distribution is as follows:

GENETICS 3: Mr 1918

disagreement of the calculations with each other may be due to the poor viability of the cultures containing less than 40 individuals. In any case, the small numbers involved render the results insignificant statistically.

But even if the second value obtained represents the true value of the coincidence, the result can be harmonized with that of the crosses involving eosin, ruby, sable, forked and cleft. For the coincidence of eosin ruby and forked cleft is a composite value; and the coincidence of eosin ruby and forked fused might be over 1.00 and of eosin ruby and fused cleft might be correspondingly less than 0.68. Thus the coincidence of eosin ruby and forked cleft might be 0.68, which is intermediate between the other two values. Why there should be a sudden drop of coincidence from more than 1.00 to 0.68 is hard to see. It might conceivably be connected with the fact that the fused cleft region is at the end of the chromosome. It is easy to see why the proportion of crossing over should be less in a given morphological distance at the end of a chromosome than in an equal morphological distance further in; for it might be supposed that the twisting of chromosomes at the end is not as tight as in regions further in. But since map distances are calculated on the basis of proportion of crossing over, any such decrease in the amount of crossing over must already have been taken into account in the calculation of the map distance; and it is hard to see why coincidence should be affected

For distances less than 46, the best data for the calculation of coincidence in the sex chromosomes are MULLER'S (1916). MULLER'S curve shows a steady rise of coincidence from 0 to a little over 1.00, as the distance increases from 0 to 45. STURTEVANT'S (1915) data, which gave a value of 2.88 for the coincidence of yellow eosin and vermilion miniature (a mean intermediate distance of about 33) are too small to be significant. MULLER'S curve also shows a fall and a second rise in coincidence when the distance rises above 45; but as MULLER himself stated, this part of the curve is not significant because the data were insufficient.

We may, therefore, conclude that in the sex chromosomes of Drosophila the coincidence rises to about 1.00 as distance increases to about 46, and that coincidence declines as distance increases further.

COINCIDENCE IN THE SECOND CHROMOSOME

In the second chromosome of Drosophila, only a few determinations of coincidence have been published. STURTEVANT (1915) obtained a co-

incidence of 0.307 in a cross involving black, purple, and curved (a total map distance of about 25.5), and a coincidence of 0.599 for a cross involving black, curved and speck (a total distance of about 55) (see figure 9). The data in both cases were too small to be significant. BRIDGES (1915) obtained a coincidence of 1.11 (first broods) and of 1.00 (second broods) for black, purple and curved. This variation of coincidence with the age of the female may not be significant; but even if these figures do not represent values significantly greater than 1.00, it is evident that there must be values of coincidence greater than 1.00 in the second chromosome. For the figures obtained represent average values, including the coincidence of adjacent regions, which is generally low (less than 1.00) when the regions are short. Therefore, the coincidence of the more widely separated portions of the regions involved is probably greater than 1.00.

Dr. BRIDGES has kindly placed at my disposal the data (as yet unpublished) of a cross involving the factors star, purple, curved and speck, whose relative positions are shown in figure 9.

star	black purple	curved	speck
		<u> </u>	1
		l	1
0	46 51	71	101
	FIGURE 9		

The data for this cross are as follows:

Table	8
TABLE	8

								TOTAL
	I	2	3	I 2	I 3	2 3	123	
192 9	1487	687	1005	601	837	135	85	6766

·

The coincidences of the various regions are as follows:

I ABLE 9								
Regions	Total length	Coincidence						
Purple curved and curved speck	53	0.4787						
Star purple and purple curved	67	1.0226						
Star purple and curved speck	99	0.9124						

These figures resemble those obtained for the first chromosome in that the coincidence first rises and then falls with increase of distance, but the figures are not consistent with the figures of the black purple curved cross, which gave a coincidence of about 1.00 for a total distance of

T +++ 1

148 ALEXANDER WEINSTEIN

about 20. These inconsistencies may be partly due to the variability of linkage in the second chromosome; they are probably also due to the fact that each of the regions involved (except black purple) is long enough to allow a considerable amount of double crossing over (which remains undetected) within it. It should be recalled in this connection that double crossing over for a given distance is more frequent in the second chromosome than in the first.

COINCIDENCE IN THE THIRD CHROMOSOME

The only published third chromosome data from which coincidence can be calculated are those of MULLER (1916). The counts are, however, very small. Some unpublished crosses made by MULLER and BRIDGES give results somewhat more extensive, but still too small to be significant. The coincidence does not rise much above 1.00, except in two cases in which it is over 2.00; but in both cases larger counts might change the results.

COINCIDENCE IN OTHER FORMS

In Primula ALTENBURG (1916) has obtained a coincidence of 0.64 (possibly, as he explained, this may represent a value as high as 1.00) for two adjacent regions whose lengths are 11.62 and 34.02 units respectively. The only other crosses hitherto reported involving more than two pairs of linked Mendelian genes at the same time have been carried out by GREGORY (1911) with Primula and PUNNETT (1913) with sweet peas. In PUNNETT's crosses it is not possible to calculate the coincidence, since he worked with an F_2 instead of a backcross; and GREGORY's results as reported give the linkage of only two pairs of factors at a time. Even so, coincidence could be calculated for GREGORY's data had he not been unable to follow one of the factors in all the plants. For given *AB*, *BC* and *AC* in any one cross, the doubles may be deduced (STURTEVANT 1914, BRIDGES 1914).

THE MECHANISM OF CROSSING OVER

It has been pointed out that during crossing over the chromosomes might be either loosely or tightly twisted about each other (MORGAN, STURTEVANT, MULLER and BRIDGES 1915, p. 64, MULLER 1916).

If crossing over occurs when the chromosomes are loosely twisted, i.e., when there are a few long loops, the low coincidence of crossing over of neighboring regions may be explained on the supposition that a twist in one region tends to prevent twisting in near-by regions. If the longer loops are more frequent than the shorter ones, coincidence will rise with increase of distance between the points at which crossing over takes place. If there is a tendency to form loops of a particular length and if loops of greater or lesser length are less frequent, coincidence will rise to a maximum for an intermediate distance corresponding to the most frequent length of loop, and will then decline. The maximum coincidence might be greater than 1.00, but it might also be less. If the most frequent length of loop is sufficiently short to allow more than one to be formed in the same chromosome, there may be more than one maximum value for coincidence corresponding to the several intermediate distances. It is evident, therefore, that on the supposition of loose twisting all the known facts of coincidence may be explained.

On the other hand, let us suppose that during crossing over the chromosomes are tightly twisted; i.e., that there are many short loops instead of a few comparatively long ones. The low coincidence of crossing over of neighboring regions is then to be explained on the hypothesis that a break in one region loosens the twisting and thus prevents breaks in neighboring regions. If, however (owing to friction, adhesion or what not), more distant regions are loosened less quickly or not at all, we should expect the coincidence of widely separated regions to rise and even to reach 1.00. MULLER has pointed out that coincidence on this scheme might also rise above 1.00. MULLER's scheme could be used to explain a decline in coincidence after it had once risen above 1.00; but it is hard to see how it could explain a decline in coincidence after it had risen to only 1.00. For since a crossing over in the eosin ruby region does not affect crossing over in the sable forked region, it can not affect the coincidence of sable forked and the region to the right of forked.

If, therefore, the determinations of coincidence in this paper are valid and comparable with each other, they seem to show that the twisting of the chromosomes during crossing over is loose; or, if it is tight, that the distance between the places of crossing over depends on other conditions than the mere tension due to the twisting.

TRIPLE CROSSING OVER

In the above calculations of coincidence, triple crossovers were excluded from the double crossover class under consideration, in spite of the fact that the triples involve crossing over in the same regions as the doubles. Of course, coincidence might be arbitrarily defined so as to exclude the triples. While it is neither necessary nor desirable to limit the definition in this way for all cases (since the word may be applied

ALEXANDER WEINSTEIN

in any sense, provided the sense in which it is applied is stated), it should be observed that triple crossing over involves conditions different from those involved in double crossing over. For in double crossing over the intermediate region remains unbroken, while in triple crossing over the intermediate region breaks. If the chromosomes are tightly twisted at this stage, the intermediate region is loosened up in triple, but not in double crossing over. If the chromosomes are loosely twisted, a double crossing over need involve only a single loop, whereas a triple crossing over necessarily involves at least two shorter loops within the same distance, as indicated in figure 10.



The coincidence as calculated in the data given (that is, omitting the triples from the double crossover class) measures the tendency of a second break to occur without the interposition of an intermediate break. If coincidence be calculated by including the triples in the double crossover class, it would measure the tendency of a break to occur without regard to whether or not the intermediate region remains intact. If the chromosomes twist loosely during crossing over, it is obvious that for the calculation of the most frequent length of loop the value of the coincidence should be obtained by omitting the triples from the double crossover class.

In the eosin ruby sable forked cross, the amount of triple crossing over is so small that its inclusion would make no appreciable difference. The value of the coincidence of eosin ruby and sable forked would be raised from 1.025 to 1.078. In the eosin ruby forked cleft cross, however, the coincidence would be appreciably raised, namely, from 0.676 to 0.878. This still leaves the coincidence of eosin ruby and forked cleft markedly less than that of eosin ruby and sable forked, but it suggests that in crosses involving a larger proportion of triple crossovers—as in the second chromosome of Drosophila—the inclusion of the triples in the double crossover class might disguise the results. For while in one cross the double crossovers might in the first cross be sufficiently more numerous than in the second to make the coincidence (calculated by including both classes) equal in one case to that in the other. This would hide the fact that coincidence in each case is a composite made up of two respectively different values.

Triple crossing over in the sex chromosome is comparatively rare, and only six cases have hitherto been observed : one by STURTEVANT (1915), one by MULLER (1916), and four by BRIDGES. If, as the variation of coincidence with distance suggests, the distance between the two breaking points of a double crossing over tends to be greater than half the length of the chromosome, the chromosome is not long enough to allow two such loops to occur and we should expect the percentage of triple crossing over to be low. It is in accordance with these facts that relatively more triple crossovers were obtained in the eosin ruby forked cleft cross (a total distance of 65.5) than in the eosin ruby sable forked cross (a total distance of 57).

It is possible to calculate the coincidence of triple crossing over in a manner similar to that of calculating coincidence of double crossing over.

In the latter case it will be recalled the formula is $\frac{x}{ab}$, where a and b

are the respective proportions of crossing over in the regions involved, and x is the proportion of double crossing over involving both regions simultaneously. In the case of triple crossing over, if a, b and c are the proportions of crossing over in the regions involved, the expected amount of triple crossing over is abc, provided the regions do not affect one another. If x is the actual proportion of triple crossing over, the coincidence is $\frac{x}{abc}$.

The coincidence of triple crossing over of eosin ruby, ruby sable, and sable forked is 0.16; the coincidence of triple crossing over of eosin ruby, ruby forked and forked cleft is 0.4858. In the second chromosome the coincidence of triple crossing over in the star purple curved speck cross is 0.4157.

The phenomenon of triple crossing over raises the question of how to calculate the coincidence of distances of which at least one is sufficiently long to allow double crossing over to occur within it. Let us suppose that in the diagram (figure II) the distance PQ is long enough to

<u>M</u>	<u>, N</u>	Р	<u>Q</u>
		FIGURE 11	

allow double crossing over to take place within it. Should these doubles be included in the calculation of coincidence? The question is really a matter of definition. Perhaps the simplest way would be to disregard entirely the double crossovers within PQ and to calculate crossing over between P and Q on the basis of the individuals which are only single crossovers in this region. This is mathematically self-consistent; for if the proportion of single crossing over within MN is a; and within PQ is b, the proportion of crossovers which are singles within MN and at the same time singles within PQ will be (on chance alone) ab.

But the interpretation of such calculations of coincidence may be misleading. For suppose that a crossing over in MN prevents crossing over near P. Then it would lower the amount of double crossing over within PQ, since the total distance within PQ available for double crossing over would be decreased. To look at it in a slightly different way, the occurrence of a crossover in MN would move a double crossover within PQfurther to the right, so that one of the breaks involved might fall to the right of Q. This would obviously increase the frequency of single crossing over within PQ at the expense of double crossing over within the same region. Consequently, the apparent coincidence of crossing over of MN and PQ might be high; but this would mean only that crossing over within MN helps *single* crossing over within PQ, for the *total* amount of crossing over within PQ would be cut down.

This suggests that it might be best to include the doubles within PQ in the calculation. We can not always in practice do this, since there may not be an intermediate factor between P and Q which can be followed. The choice still remains, however, of counting each double as one crossing over or as two. The matter is again a question of how we choose to define coincidence.

The shortest distance in the sex chromosome within which a double crossing over has been observed to occur is 13.5. Hence, the considerations just mentioned do not call into question the validity of the calculations of coincidence of eosin ruby and forked cleft or of eosin ruby and sable forked. For the eosin ruby and forked cleft distances are too short to allow double crossing over to occur within them; and while the sable forked distance is just sufficiently long, the frequency of such double crossing over is so small (only one case has been observed in all the Drosophila work) that the result would not be appreciably affected. The ruby lozenge distance, which is 16 units long, is also too short to be appreciably affected.

In the second and third chromosomes of Drosophila the shortest distance within which double crossing over has been observed to occur is shorter than for the sex chromosome. Moreover, as has been pointed out, the distances in the second chromosome for which coincidence has been calculated are so long as to allow a great amount of double crossing over within them. Consequently, the interpretation of these data should be attended with caution.

MAXIMUM AND MINIMUM COINCIDENCE

It may be interesting to compare the observed values of coincidence with the maximum values mathematically possible under the circumstances. The latter values can be calculated as follows:

If a and b are respectively the lengths of (proportions of crossing over within) the regions under consideration, the amount of double crossing over involving both regions simultaneously is abx, where xis the coincidence. Now, the maximum number of double crossings over will occur when every crossing over in one region is also a crossing over in the other. That is, when coincidence is at a maximum,

$$abx = a, x = \frac{1}{b}$$
 where $a < b$
 $abx = a = b, x = \frac{1}{b} = \frac{1}{a}$, where $a = b$

That is, the greatest possible coincidence of two regions is the reciprocal of the frequency of crossing over in the greater region (if the regions are unequal), or of either region (if the regions are equal).

The greatest significant coincidence observed in the sex chromosome was 1.025, in the case of eosin ruby and sable forked. The maximum possible coincidence in this case is $\frac{I}{0.1355} = 7.3776$. The coincidence actually obtained falls far short of this.

The formula for maximum possible coincidence can be extended to any case of multiple crossing over. For example, in the case of triple crossing over, where a, b and c are the distances involved and x the co-

incidence, abcx = a, $x = \frac{1}{bc}$ where a is the shortest distance involved.

If a equals b,
$$x = \frac{\mathbf{I}}{bc} = \frac{\mathbf{I}}{ac}$$
. If $a = b = c$, $x = \frac{\mathbf{I}}{ab} = \frac{\mathbf{I}}{bc} = \frac{\mathbf{I}}{ac}$.

The minimum possible value of coincidence is obviously 0; but this is true only when the sum of the proportions of single crossing over within each of the regions involved is not greater than 1.00. For it is evident that if this sum is greater than 1.00, some crossovers in one

ALEXANDER WEINSTEIN

region must also be crossovers in the other and the coincidence must be greater than o.

THE DISTANCE BETWEEN THE BREAKING POINTS IN DOUBLE CROSSING OVER

Given the factors M N P Q in the order mentioned (see figure 12)

the average length of a double crossover involving the regions MN and PQ is $NP + \frac{MN}{2} + \frac{PQ}{2}$. For every such crossover necessarily includes NP, and it has an equal chance of involving in addition any length from o to the entire distance NM on one side, and o to PQ on the other. It would on the average, therefore, involve half of NM and half of PQ. This value is obviously the mean between the longest and the shortest possible distances, or it is the distance from the midpoint of NM to the mid-point of PQ. Where only three factors are involved, NP is o; hence the value of the mean intermediate distance becomes $\frac{MN}{2} + \frac{NQ}{2} = \frac{MQ}{2}$, or one-half the total distance.

This proof rests (as has been mentioned) on the assumptions that a break is just as likely to occur at any point as at any other, and that the distance between breaks (length of a double crossing over) is no more likely to be of any particular length than would occur as the result of pure chance. The first of these assumptions is justified, since the term distance is used to represent the frequency of crossing over and not necessarily morphological distance. The second assumption, however, is not true in general; for, as we have seen, coincidence varies with the intermediate distance, and only for some distances (where coincidence is 1.00) are double crossovers as frequent as they would be on pure chance. Hence, in any case the true average value will be a weighted mean. Since we do not know accurately how coincidence varies with short increments of distance, we cannot judge just how accurate the formula is in any But if variation in coincidence is gradual, the formula must be case. a good index of the mean intermediate distance when MN and PQ are short.

Double crossing over within short distances is very rare, and in the sex chromosome has never been observed to occur within distances shorter than 13.5. If, therefore, in any sex-linked cross in Drosophila the intermediate distance is less than 13.5, allowance can be made in the formula for the mean intermediate distance for the absence of the

shorter double crossovers by substituting 13.5 for NP provided that MQ > 13.5.

When the distance NP is longer than 13.5, double crossing over may occur within it, and if no factors in this region are followed the double crossing over can not be observed. Hence the apparent amount of crossing over in the intermediate region will be less than the true value. For example, the value obtained for the distance between ruby and sable in the eosin ruby sable forked cross was 0.333; and the value obtained for the distance between ruby and forked in the eosin ruby forked cleft cross was 0.415. The map values for these regions are respectively 36.5 and 51.5.

It is, of course, possible, though unlikely, that the amount of undetected double crossing over within the ruby sable region in the first cross was sufficiently greater than the amount of double crossing over within the ruby forked region in the second cross to make the ruby sable region in the first case greater than the ruby forked region in the second case. This is further suggested as a possibility by the fact that the crossing over in the eosin ruby region is slightly greater in the first case (0.0668) than in the second (0.0560). And this might be held to explain why the coincidence of eosin ruby and forked cleft is lower than that of eosin ruby and sable forked. But even if we increase the map value of the ruby sable distance in the ratio of $\frac{0.0668}{0.0560}$, the distance will

be only 41.61. This is still considerably less than the map distance of ruby forked and is almost exactly equal to the apparent length of the latter distance in the eosin ruby forked cleft cross. Since the true length in the latter case must have been greater than 41.61 because of the occurrence of (unobserved) double crossing over within it, the suggestion that the ruby sable region was genetically longer cannot be considered probable, though it remains as a rather remote possibility.

To dispose of this possibility absolutely it would be necessary either to follow enough factors in the intermediate region so that no double crossing over remains unobserved, or to make a cross involving simultaneously all the regions whose coincidences are sought The former method has the disadvantage that a great number of mutant factors tends to cause differential non-viability and that it is not always feasible to obtain properly spaced factors which can be worked together. The second method was actually tried by making crosses involving simultaneously eosin, ruby, sable, forked and cleft. But it was found that sable cleft flies were

ALEXANDER WEINSTEIN

almost always non-viable and the cross was abandoned. The method was, however, successfully used in the cross involving simultaneously eosin, ruby, lozenge, forked and cleft. Here the ruby forked distance is necessarily longer than the lozenge forked distance, since the former consists of the latter plus the ruby lozenge distance. It will be recalled that the results of this cross, while not numerically great, were in accordance with those of the two main crosses.

MUTATIONS OBSERVED

Several mutations were observed in the course of this work.

I. Yellow body color. Three yellow males appeared in a cross of a female carrying the factors eosin ruby forked in one sex chromosome and the factor fused in the other, by a bar male. The other offspring fell into the expected classes. Of the three yellow flies, one was also eosin fused, a second eosin ruby fused, and the third eosin forked; so that they also (except for the yellow character) fell into expected classes. This, together with the fact that the amount of crossing over between yellow and eosin is only one percent, and that no other yellow eosin flies were then, so far as known, in existence, makes it quite unlikely that the three yellow flies were the result of contamination. The yellow factor must therefore have arisen by mutation in the sex chromosome of the mother, for the mutants were all males and did not arise by non-disjunction.

The new yellow was ascertained to be sex-linked and recessive to wild, like the old yellow. When it was mated to the old yellow the daughters produced were yellow; hence the two factors must be the same.

2. Achete. The laboratory stock of yellow was discovered to be pure for a factor causing a reduction in number, and sometimes a total absence of, the dorso-central bristles. This factor, termed achete, is a sex-linked recessive. No crossovers between it and yellow were observed in over 200 flies; hence it is either closely linked to yellow or an effect of the yellow factor itself. If the latter turns out to be true, the yellow locus may furnish a case of quadruple allelomorphism, for another yellow stock (containing also white eyes) and the yellow which arose independently (as reported above) have the normal number of dorso-central bristles (four), and there are besides two other allelomorphs, spot and normal.

3. Lethal. Two females in the eosin ruby sable forked cross gave lethal ratios. The lethal in each case is about one unit from yellow; in one case it is known to be to the left of yellow. It is, therefore, the



FIGURE 13 .- Fly with inflated wings.

farthest to the left of any factor known in the sex chromosome. Since the females were sisters, the same factor is probably involved in both cases.

4. Inflated. In several of the cultures of the eosin ruby forked cleft cross there appeared flies whose wings were inflated (figure 13). As the flies grow older the wings collapse and look blistered. This variation was ascertained to be sex-linked and to be located about I or 2 units to the left of forked, in approximately the same locus as the factor for rudimentary wing. But the two factors are not allelomorphic to each other, since the F_1 females of the cross between them are long-winged.

In several of the cases reported above, more than one individual displaying the same mutant character appeared in the same brood. It seems unlikely that the character arose independently in each individual; more probably the individuals in each case were derived from a single mutated germ cell. If this is true, the yellow mutation must have occurred at least before the next to the last oögonial division, since it arose in the female; the lethal factor must have arisen at least before the first maturation division if it occurred in the mother of the females tested, but it may have originated between the first and second maturation divisions if it arose in the father.

GENETICS 3: Mr 1918

SUMMARY

It has been known that the coincidence of crossing over of two regions increases in general as the distance between them increases. The evidence presented in this paper indicates that, for the sex chromosome of *Drosophila melanogaster*, when the intermediate region reaches a value of about 46, coincidence is approximately 1.00; and as the intermediate distance increases still further, coincidence decreases again. In other words, crossing over in one region of the chromosome interferes with crossing over in neighboring regions. But this influence decreases as the distance between the regions increases, until when the distance is about 46 units the interference has practically disappeared. For regions more than 46 units apart, interference reappears again.

It is pointed out that if the data presented are statistically significant, either the twisting of the chromosomes during the process of crossing over is loose, or the distance between the places of crossing over in the chromosome is dependent on other conditions than the mere tension due to the twisting.

I wish to thank Professor T. H. MORGAN and Dr. H. J. MULLER, Dr. C. B. BRIDGES and Dr. A. H. STURTEVANT for helpful suggestions made in the course of this work. I wish also to thank Dr. RAYMOND PEARL, Dr. J. A. HARRIS, and Mr. J. W. GOWEN for help with the statistical aspects of the problem.

LITERATURE CITED

ALTENBURG, E., 1916 Linkage in Primula sinensis. Genetics 1: 354-366.

- BRIDGES, C. B., 1914 The chromosome hypothesis of linkage applied to sweet peas and Primula. Amer. Nat. 48: 524-534.
 - 1915 A linkage variation in Drosophila. Jour. Exp. Zoöl. 19: 1-21.
 - 1916 Non-disjunction as proof of the chromosome theory of heredity. Genetics1: 1-52, 107-163.
- ELDERTON, W. P., 1902 Tables for testing the goodness of fit of theory to observations. Biometrika 1: 155-163.
- GREGORY, R. P., 1911 a Experiments with Primula sinensis. Jour. Genetics 1: 73-132. 1911 b On gametic coupling and repulsion in Primula sinensis. Proc. Roy. Soc. 84: 12-15.
- JANSSENS, F. A., 1909 La theorie de la chiasmatypie. La Cellule 25: 387-411.
- MORGAN, T. H., 1910 An attempt to analyze the constitution of the chromosomes on the basis of sex-limited inheritance in Drosophila. Jour. Exp. Zoöl. 11: 365-414.

1912 The heredity of body color in Drosophila. Jour. Exp. Zoöl. 13: 27-46.

- MORGAN, T. H., and BRIDGES, C. B., 1916 Sex-linked inheritance in Drosophila. Carnegie Institution of Washington, Publ. 237. 88 pp.
- MORGAN, T. H., STURTEVANT, A. H., MULLER, H. J., and BRIDGES, C. B., 1915 The mechanism of Mendelian heredity. xiii + 262 pp. New York: Henry Holt.

159

- MULLER, H. J., 1916 The mechanism of crossing over. Amer. Nat. 50: 193-221, 284-305, 350-366, 421-434.
- PEARL, R., 1917 The probable error of a Mendelian class frequency. Amer. Nat. 51: 144-156.

PEARL, R., and MINER, J. R., 1914 A table for estimating the probable significance of statistical constants. Ann. Rep. Maine Agr. Exp. Sta. pp. 85-88.

- PEARSON, K., 1911 On the probability that two independent distributions of frequency are really samples from the same population. Biometrika 8: 250-254.
 - 1914 Tables for statisticians and biometricians. pp. lxxxiii + 143. Cambridge Univ. Press.

1917 The probable error of a Mendelian class frequency. Biometrika 11: 429-432.

PUNNETT, R. C., 1913 Reduplication series in sweet peas. Jour. Genetics 3: 77-103.

- STURTEVANT, A. H., 1913 The linear arrangement of six sex-linked factors in Drosophila. Jour. Exp. Zoöl. 14: 43-59.
 - 1914 The reduplication hypothesis as applied to Drosophila. Amer. Nat. 48: 535-549.

1915 The behavior of the chromosomes as studied through linkage. Zeitschr. f. ind. Abst. u. Vererb. 13: 234-287.

- WRIGHT, S., 1917 On the probable error of Mendelian class frequencies. Amer. Nat. 51: 373-375.
- YULE, G. U., 1911 An introduction to the theory of statistics. pp. xiii + 376. London: Charles Griffin & Co.

NOTE ADDED IN THE PAGE PROOFS.

At the bottom of page 145 the value of the coincidence of eosin ruby and forked fused (for all the broods, including those containing less than o males) should be 0.8572 instead of 1.1144 as given. (The latter value represents the coincidence calculated by including the triples in the double crossover class.) The suggestion made on page 146 in connection with the value 1.1144 is rendered unnecessary, since the correct value agrees with the expectation and is consistent with the other results obtained.

TABLE 10

Individual culture counts of crosses involving easin (w^e), ruby (r_b), sable (s), and forked (f).

	Coincidence	of		0	0	1.7714 0	4.1111	4.1667	2.5550	°, 0	0	o c	0	0.9667	I.2436	1.9286	0.7390	00	1.6600	0	0	1.8000	2.2778	0	0	•	
		Total males		68	114	02	74	50	38	72	8 i	55 52 52	,8 8	58	97	54	1.	28	83	101	62	57 420	82	49	6	85	
		1 2 3	w ^e r _b s f								I															I	
		3	2 _d 7		ŝ	N		0	4		,	N H						אמ		I	H	I	I			ŝ	
·			wef	I	H	н н		,	00	0	3	H 0	I	ŝ	4		H			~	I	H	61			I	
		3	5					I						I	I	F	H										
		I	$w^{\rm e}r_{\rm h}f$			H	N	•							н				I			I	I				
ي الح		1	*+	ц		щ		۲	-	I			I	I	0		H			N		I	I			H	
3	ıles	1	$w^{e}r_{b}sf$			н									01		I	I								₩	
(I)	Ma	3	$r_{\rm b}f$	N	0	4 -	с С	ς, τ	4 ں	• 4	I	I	ر ر ا	Ŋ	Ŋ	H.	ທຸ	∨ 4	4		4	I	I	4	I	v	
			w ^e s	2	œ	- ~	H	•	4 (1	I	I	01	3	I	F	2	4,	-	ς Li	7		I	4	10	T	3	
			r _b sf	II	14	10) OI	нv	2 4	- v	~	۲ <u>م</u>	9	ŝ	II	II	6 ¦	19 19	7	15	10	v,	16	61	v	IO	
		7	w ^e	18	19	۲ . 6	15	13	11	17	12	51	01	12	21	II	1	191	14	15	6	13	13	0,	×	18	
			sf		I		3	H	H	3	4	∼	2	ю	4	N	с о (m 01	1	4	3	H	н	H	-		
		_	w ^e r,	ę	0	ω 4	. 01	н (ч н	I	4	H		н	3	ৰ	91	- 0	ŝ	~	N	<i>с</i> о	-	H	N	61	
		1	۴ ه	22	29	12 12	, K	51	38	26	17	ວຼ∝	18	13	19	12	10 10	10	8	22	12	13	21	17	I.S.	22	d type
			w ^e sf	20	28	13	16	13	12	12	ē,	<u>0</u> 2	14	15	23	OI	13	19	61	52	21	13	8	13	~	81	tes wil
		Females		ÕП	×	₹&	001	8 5	16	104	8	7 04	22	8	00 G		48	88	SoI	46	74	75	16	7	2	88	bol + indica
		Culture No.		125	120	127	300	225	227	228	241	<u>5</u> 8	203	224	220	231	233	2,36	230	240	245	246	128	194	202	281	* The sym

GENETICS 3: 160 Mr 1918

TABLE IO (continued) (1) $\frac{w^{\bullet} sf}{r_{h}}$ (continued)

	Coincidence of	$w'r_b$ and sf	3.8750 2.0444 2.0444 1.5987 0 0 0 1.3036 0 1.3036 0 0 1.3538 1.8750 1.9800 1.9800 1.9800 1.5952 1.5952 1.55588 1.5558 1.55588 1.55588 1.55588 1.55588 1.5558	
	Total males		8%3 <i>6717</i> 88%58888888788878888888888888888888888	3957
	3	f		н
		werbs	·	I
	2 3	S ^{iq} 1	а аа ннаюаны ы ю	46
		$w^{e}f$	H 10 H 10 0 H 10 0 H 4 10 H H	8
	5	5	H H HH Ø HHØHH	61
	<u> </u>	$w^{e}r_{b}f$	ны о на на на на	22
		+	H 0 0 0 H 4H H HH	31
lles		werbsf	HØ HH Ø H4ØH HHH	25
Ma	3	rbf	4 4 H WW 4 Ø 0 W 0 0 4 0 W Ø 4 4 4 H W K Ø K Ø	192
		w ^e s	<u>ини40000 но044004000 440н000</u>	145
	2	r _b sf	110872510010010825875747829255 229255	488
		200 ^e	8005135135135855555555555555555555555555	780
		sf	- のてののちてのここののことのアナジの下ジのの	102
	-	wer,	550000 00 15500011 0415140	105
		t _b	8 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2	1030
		west	188354599144888888888888888	904
	Females		5 5 8 8 8 9 8 8 8 8 8 8 8 8 8 8 8 8 8 8	4726
	Culture No.		851 152 262 265 265 265 265 265 265 265 265 2	Totals

GENETICS 3: 161 Mr 1918

TABLE IO (continued)

werb f	s
3	Ì

Coincidence	of w ^e r. and sf	4	1.5962	1.4808	2.4007		5.0000	0	0	1.9333	1.3333	0	0	2.1714	0.1480	1.7374	0	5.9000	0.8011	2.0857	0	1.3810	0	0.9630	110.1	1.0833	0	0
T_{0}	u ulai males		83	22	62	લ્ક	38	75	143	116	120	107	80 20	20	107	ଛିଁ	04	59	93	73	74	87	120	52	58	211	26	58
	<u>1 2 3</u>	$w^{e} r_{b} sf$																										
	23	$w^e r_b sf +$	2 I	II	I I	• •	4 H	4		I	I	I	1	I	I	I		I	4	I	ŝ	61	4	H	H	1	ы 19	2
	<u></u>	r _b	N	N			-	•						H	н	н		н						I		I		
	н	w ^e sf			10					I	I			61		I			I	П		ы			I	I		
		r _b s		I		п								I					н	I	0	н	0	0			I	
es	1	$w^{e}f$	10	I	I	ı	-	0		н	I	I					I				I			I				
Mai	<u>د</u>	sf	~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~	Ŋ	vu/	0,	- ر×	•	Ŋ	7	0	ŝ		9	9	(1	ę		×	0	6	ŝ	ę	4	0	7	×	F
		w ^e r _h	ъ N	4	2	4 (N	• •	8	S	7	0	e S	ŝ	6	9	ŝ	ŝ	ŝ	0	4	ъ N	7		ъ	13	ŝ	3
	1	f	~	10	10	41	~2	17	21	6	8	18	14	15	12	13	15	14	17	0	I	14	24	~	12	22	11	01
		20 er bs	∞	15	4	11	11		17	12	11	12	6	01	10	II	ν	ŝ	∞	13	∞	17	17	9	9	14	14	12
		$r_{\rm b}f$	4	I	I	0 F	- 6	ب د	N	I		H	c,	e		4		I	ę	Ι	0	e	~	ę	I	S	n	н
	-	w ^e s		3		0 I	n .	Q	0	0	4	ę			3	3			I	0	I	ę	61	0	0	I	ъ	H×
		S	23	11	53	52	1	81	40	28	36	54	23	12	31	21	14	21	19	22	17	16	25	10	13	21	23	10
		$w^{e}r_{b}f$	36	22	54	12	5 S	<u>8</u>	5 5	49	33	4	30	8	33	23	53	15	26	18	16	61	58 78	15	15	27	21	12
- F	l'emales		<u>8</u>	011	108	134	201	87	155	149	147	134	105	118	122	103	20	IOS	×	104	62	85	126	81	20	135	135	<u>8</u>
Culture	No.		213	214	215	210	210	250	302	303	306	310	323	223	305	208	210	211	212	220	221	222	251	252	253	254	250	257

GENETICS 3: 162 Mr 1918

TABLE IO (continued)

$\frac{w^{e}r_{b}}{s}f$	
(3)	

Coincidence	of mer and cf	te nue qu m	0.8750	1.9444	0	0.8407	3.2000	0	I.5850	2.4103	0	0	0	0	。 。	1.8286	0	0	2.1901	•	0	1.2750	1.3500	0.4328	0.5926	1/01.1	1.3175	3.7143	0
Total	u utat males		91	20	29	153	112	80 v	82	2	124	22	83	50	103	64	106	16	112	48	74	51	54	74	8	62	83	25	58
	123	w ^e r _b sf																						I					
	2 3	w ^e r _b sf +	3 I	0 0	I 2	I 3	I I	2 1	IS	. 3	I			2 I	64		II	I 2	I 3		I	I 3	0	0	1 6	I	2		I
	<u>6</u>	r _b		I		Π	0		I	I						н			I	-		н	н	I		н	н	I	
	-1	w ^e sf	н				I	_			_														I				
		r _b S	-			•				I			I					I			н			I	н				
les		$w^{e}f$			I			0								H				н	I	н		0		н			н
Mal	3	sf	4	~ ~1	4	II	4	ŝ	4	Ś	8	4	4	4	7	4	e	N	10	0	4	п	4	9	9	ę	0	Ŋ	с
		wer,	4	4	()	10	12	9	5	4	01	9	9	61	4	0	œ	ъ	0	I	ę	0	н	6	4	61	6 0	H	4
		f	13	II	14	35	20	16	4	19	18	4	II	7	16	II	18	01	20	~	81	6	7	IJ	17	II	12	7	II
		20 ° r b S	13	12	II	23	18	12	13	4	01	ŝ	ø	∞	II	9	16	0	18	3	×	ഗ	4	H	14	∞	10 I	ŝ	ŝ
		r _b f	4	I		Ι	0	I	I		ę	I	0	I	I	I	4	9	I	10	4	0	e	1	4	3	0	н	ŝ
	-	w ^e s	10	H	ę	U)		I	I	I	I			I	0	0	19	ę	I	н	H	I	н	ę	ę	0	4		I
	-	5	24	15	23	32	5	6	53	33	31	24	30	12	43	13	82	ų	50	S	14	13	61	01	13	15	20	6 1	8
		$w^e r_b f$	21	18	17	41	27	29	22	33	4	33	20	12	17	21	25	27	8	21	19	12	12	12	50	14	20	15	01
L	Females		93	64	83	160	149	120	8	112	116	22	8	52	103	85	102	114	105	8	20	8	\$	75	011	71	82	78	82
Culture.	No.		259	260	261	301	309	311	312	314	315	317	318	319	320	321	322	324	325	310	282	283	284	285	286	287	292	357	358

GENETICS 3: 163 Mr 1918

•

TABLE 10 (continued) (2) $\frac{w^{r_{p}} f}{s}$ (continued)

Coincidence	of w ^e r. and sf		2.90000	0	0	0	0/0	0	1.5667	0	0	0	0	2.2941	•	•	3.1818	1.7400	0	0	0	0.8333	0	6.3333	0	0	0	
	1 otal males		58	8	92	4 3	120	107	94	131	62	53	112	52	92	59	2	87	20	88	84	98	77	20	æ	8	84	6623
	23	r _b sf															н	~										I
	нI	ร้																										н
	2 3	+	0	I	N	н	0 F	((N	I	4			I		3	I	0	F	e	I			I	н	н	I		103
		w ^e r _b S]	-	01		H	~	°	I	3			(1)		I		61	1		61						3		85
	3	* ^b	-											ŝ				H				H		I				32
	-1	w ^e sf					۰	1	I								61	0										23
		r _b S		•		I	۰ ر	1	I											I								25
les	1	$w^{e}f$		I			-	1				I						I		I		н	I					31
Maj	3	sf	3	e	e	H	44) er	, m	6	н	н	~	3	~	0	I	Ŋ	<i>.</i>	I	0	Ś	I	e	ŝ	4	~	323
		w ^e r,	3	40	×	9	<u>ن</u> ن		o o	3	ŝ	Ŋ	×	II	e	3	ς	4	2	×	ю,	9	v	~	3	ŝ	4	390
	1	f	6	15	9 <u>0</u>	× ·	4 ç	2 8 1	18	22	01	14	22	01	13	01	10	9	~	17	20	12	20	12	22	15	12	078
	10	werbs	9	6	6	ונע	чç	21	12	17	0	0	×	v	6	0	12	11	ъ	4	4	4	∞	~	~	∞	3	734 I
		$r_{\rm b}f$			4	~~~		101	I	I	I	n	3		9	N		3	F	0		0	I			ŝ	Ĩ	150
	-	w ^e s	-	ъ		I	ý)	61	ŝ			61	I	I	6	I	ę	3	ŝ	ŝ	ы			H		H	133
		S	13	31	Ş	0 I	រូ ខ្	26	5	41	23	II	ନ୍ଧ	6	21	16	22	61	61	30	23	~	19	21	8	24	20	693
		$w^{e}r_{b}f$	61	5 8	24	II	II 6	55	21	5 8	8	16	3 0	01	88	21	14	8	22	16	27	20	21	24	21	27	27	1821
ŗ	Females		22	107	001	27	78	911	145	134	83	64	112	8	93	26	11	95	III	83	III	26	103	6	123	611	141	8022
	Culture No.		359	395	405	200	352	331	332	334	337	338	339	340	341	342	343	345	346	347	348	349	350	353	355	300	368	Totals

GENETICS 3: 164 Mr 1918

•

TABLE IO (continued)

	Coincidence	of mer and ef	fe nup q 1 m	0.5229	2.6667	0	ວັ	1.6250	I.9286	1.2857	1.8846	3.3333	0	٥`	1.6000	0/0	•	1.4833	0.8095	•	5.0000	
		Total males		&	8	53	63	117 81	24	108	<i>&</i> ,	8	4 <u>6</u>	46 6	48	42	45	£	102	19	00	1349
		<u>1 2 3</u>	$w^{e}r_{b}sf +$								I						ш [°]					ΙΙ
		2 3	w ^e r _b sf	6 6	2	I	ę	I	I	I I	s,	33	I	2			N	3 I	I	I	I	23 20
		3.	sf	I	-			I	I		н	I							I			
		I	wer,		I				F	I	3	I			H,			I			н	01
÷.		5	f					Ħ.		N							01		2	~		6
r, S	ales		Werbs	10							1											5
(3) 2	M	3	٩ ^٣		4	4	H I	× ×		9	ŝ	61	3	п		H	61	ъ	~	17		51
Ū.			w ^e sf	4	<u>س</u>	-	01	н V	<i>n</i>	~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~	ŝ	0	H	3	v	-		~	ŝ	I	H	 22
		2	1.bS	7	17	6	IO	6I 1	່	12	II	I	4	e		ŝ		12	12	H.		142
			$w^{e}f$	2	15	13	6	<u>6</u> ;	9 Y	. 6I	24	12	<u> </u>	∞		6	II	12	6I	17	33	272
			bf s		I	10	3	01 F			0						-	-	4	2	•	5 25
			f wer				<u>س</u>				•	~	- 		- -							4
			•S 1	2	ล ถ	51 O	5	Эл	- 11 - 00	80 (43)	ъ б	э 9	4	6 2]	2	0 2	ਸ 8	5 3	50 10 10	ы С	ĩ	o4 38,
		s:	n				-	<u> </u>	•		н —		н ——					Η.	~			Ř
		Female		108	88	<u></u>	67	061 092	812	141	72	I02	72	87	IOI ·	72	75	108 108	152	102	1/1	1908
•		Culture No.		328	329	329 a	329b	377	384 384	386	390	393a	369	371	385	391	<u>30</u> 3c	380	382	370	379	Totals

GRNETICS 3: 165 Mr 1918

TABLE IO (continued)

`

(4) $\frac{w^{e}r_{b}sf}{w}$

Coincidence	of	w ^r _b and sf	1.7308 0	00	3.5000	0 2.5000	
	Total males	1114105	96 87	47	0 1 04	47 45	458
	1 2 3	wes ruf					
an and an a star with a subscription of the su	2 3	w ^e r _b f s	<i>a</i> w	•	-	I	6 I
	<u>1</u> 3	$w^{e}f r_{b}s$	I		п	I	3
es	1 2	w ^e sf r _b		I I			II
Mal	3	$w^{e}r_{b}s f$	4 5 13 7 7	ب میں د	- 0	T Q	13 44
	2	wer, sf	17 17 15 13 8 6	1 N G	2 ~ ~	000 000	76 68
	-	$w^{e} r_{h}sf$	4 1 3	ς, ⊢	- н	2 I I	13 5
		$w^{e}r_{b}sf$ +	18 22 19 25 18	7	5 23 5	9 6 13	70 157
	Females	-	116 19 19	:6:1	82	53	642
	Culture		148 151 147	140	150	153 154	Totals

GENETICS 3: 166 Mr 1918

TABLE II

Individual culture counts of crosses involving easin (w°), ruby (r_{b}), forked (f) and cleft (c_{t}).

$\frac{w^{e}r_{b}f}{c_{t}}$
(1)

Coincidence	$\int_{T}^{\infty} e_{F_{r}}$ and $f_{C_{r}}$		0	0	0	0	0.0000	0	1.8250	0	0	•	3.8214	0	•	•	•	•	1.6667	0	0	0/0	0	0	5.5625	0	0	0	0
E	Total males		134	62	87	81	15	2°	73	99	54	73	201	65	107	122	63	19	20	82	64	11	72	6	æ	20	58	73	22
	<u>1 2 3</u>	w ^e fce r _b																											
	2 3	$w^{\rm e}r_{\rm b}$ fc _t	I I		3 I	0	÷	•	I	ΙΙ		I			I	I I	Ι		I I	II	I	I	2			I	Ι	I	I
	<u>1 3</u>	we rofc.					I		7				I						н						I			•	-
2	I 2	$v^{e}f r_{b}c_{f}$	F		I			н	I 2		2 I		61		I I	~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~				3 I	I		-	I				I	I I I
Male	3	$v^e r_b f c_t + \frac{1}{2}$	I 4	I	5 I	ы 1 1 1 1 1 1	3 1	I	3 4	-	3	I I	I S	3 3	I 4	2 7	3 2	3 I	I 2	e	I 3	ω.	г	5	и 1 9	1 3	0	0	1 I
	7	ber f z	25) 12	61 2	52		2 IG	LI 0	5 I8	II 8	22	28	3 17	24	33	15	01	0I Q	5 14	0I	6	8	7 21	ZI 2	22	22	5 I3	101 0
		rbf wer	3	Ξ'	3 II	н ; н ;			3	п			N I		i I	H I	E N		е С	1 1 1	i 8		3 I(Ξ' 	0	0	1 7	3 - 0
	н	w ^e c _t	10	I			0	I		ŝ	н	0		I	<i>ლ</i>	17		I	01	<i>ლ</i>	I		I	I	I		I		5
		of cr	37	23	21	83	-1 OI	I0	IO	14	II	12	IS.	13	33	22	II	14	22	23	21	21	18	14	-18 	с С	9	15	21
	emales	wer	160 37	95 23	98 21	108 20 20 20	9/ 57	80	106	79 21	82 9	81 25	175 34	84 IE	146 31	119 41	63 IG	58	106 20	96 I5	83	87 I6	91 22	80 21	96 30	120 20	87 IS	89 23	02 10
	No.		4	ر ي .	7	6 <u>;</u>	10	16	17	18	19	50	21	22	24	25	32	33	34	35	36	54	55	56	57	\$	65	73	74

GENETICS 3: 167 Mr 1918

TABLE II (continued)

(1) $\frac{w^* r v f}{c_1}$ (continued)

Coincidence	of $w^e r_b$ and fc ,	-	0	0	0	0	2.4000	0	0	0	0	0	0	4.7333	0	0	0/0	. 0	0	0/0	0	0	0	0.5000	0	1.5312	1.5515	3.0357	0	5.9000
E	1 otal males		102	16	54	82	<u>6</u> 0	6	67	55	66	60	63	71	56	40	8 5	53	50	59	20	02	8	84	85	<u>8</u> 6	128	85	8	59
	I 2 3	wefce rb																												
	2 3	fcr	I		I	I			н	I	I	I							H		-			61	61	3	I	I		
		t werb	н	H	0	3	01		I	I	H	H	I		I		3			I	64		3	ŝ	4	4	0	0	I	
	3	rufe												I													61	I		8
		æ					-								_									н		I				
1	10	rbcr			I		0	ę			I								I						F	I	I			
ales		w ^e f		N	0	H	I			н				H				-	0				0	4			9			••
W	3	$w^{\epsilon}r_{b}fc_{t}$ +	5 3	Ι	I	0	0	3 I	I 3	II	I 4	3	2 I	5		I	1 3	1 2	I	3 2	I I	4	6	6 7	3	3 5	2 4	2 I	60	3
		f	6I	53	10	16	7	14	10	16	14	12	16	18	8	13	II	11	~	14	20	II	26	16	15	17	24	13	24	20
	8	werbcr	14	18	II	18	14	16	~	9	21	12	4	œ	01	II	н	IO	12	~	15	2	17	II	12	14	21	13	15 2	×
		$r_{ m b}f$	I	0	0	0		ę			ы		0	I				I			H	6		I	0	61	4	I	I	N
	п	$w^e c_f$			0	4	I	ŝ	I		I	I	0		I	I		I	Ţ				0	0	3		ы	61	I	
		Cr	25	15	0	14	13	50	81	14	27	12	15	18	12	7	13	10	13	II	~	6	18	13	15	18	33	26	15	~
		$w^{e}r_{b}f$	33	8	13	21	17	27	24	14	20	25	ଷ	20	12	2	20	16	12	21	6	71	20	16	27	30	କ୍ଷ	23	8	17
•.	Females		211	103	83	104	88	124	8	82	00 I	55	8	78	84	37	20	70	8	22	62	85	102	93	106	8	129	<u>9</u>	χ,	8
Cuttures	No.		75	×	II	69	20	23	37	30	40	48	49	51	52	53	50	62	63	49	22	78	81	82	84	88	48	88	ŝ	8

GENETICS 3: 168 Mr 1918

TABLE II (continued)

 $(\mathbf{I}) \frac{w^e r_b f}{c_I}$ (continued)

Coincidence of	tal Coincidence of les $w^* r_b$ and fc		00	2.8007	00	1.1477	•	00	0 0	00	1.3750	0 0	00	0	0	4.0000	0	0	0	0	0	0	7.5455 0	
Total	males	123	ଟ୍ରମ୍ବର	8,4 8,4	82 92	101	52	53 81	6 <u>5</u>	2 2	8	12	5/ 12	20	55	72	8	ς δ	112	III	136	1.58	83 124	
1 2 3	efce ro	1	I	-									F	I										
~	fcr a	010	n w	0	010	30	I	н		N	I	I	H	ч		· · ·	4		F	I	I	ю —		
~	rbfcr 20°rb	3	н	N		1 2	2	04	. 1			•					I		<u>л</u>		N	0	0 H	-
-	ze •			H							I													
1	$\frac{1}{w^{e}f} r_{b}c_{1}$	5	5 7 7	4 1 1	п 1	I I I		I			I I	н (4 14	I 3				I I	N		0	N		
Male	$\frac{1}{e^{r_bfc_t}}$	3. 2	5 C	1 4	, , H	04 00	II	4 3	н () м (ч н ч н	4	н с с	4 با ن	6 I	а н н	- 6	1 3	5	3 4	დ. 4	40 m		4 0 2 m 	
~	bcr f 2	20	282 282	53 53 23	5 5 5 7 7 7 7	181	11 6	3 11		0 00 1 13 2 13	5 II	13	11	. 0		4 01 4 01	10 10	0 15	4 2I	22 25	3 23	27	31 - 8	ES,
-	rbf wer	8 8 8 8 8 8	N (N N (N	м 	н с н		2	н н 	а ,	-	0	а н н (N	4	01		, <u> </u>	1	3	6 6	-0 	4	 	by BRIDGI
н	wect	3	- 01	- 9	ς, τ	- 0	г	H 4	· H +	ч н	ŝ	0 F	- 4	. 01	<u>н</u> (- ري	4		9	н	H	m	п 	ies made
	verbf cr	27 29 21 29	34 53 34 23	23 35 32 32	18 18 20	5 7 5 7 5 7	16 8	10 16 14	01 01	19 14 10 17	20 17	22 10	10 16 23	20 13	25 11 25	20 20 20 20	I4 20	27 20	33 20	31 22	43 27	42 32	27 19 30 23	from cross
Females -	6	112 8r	ده 142	156	211	001	<u>7</u> 9	¥.8	104	8 %	202	8 Y	28 28	81	82	5.8	62	113	124	133	601	152	05 135	2388-2505 are 1
Culture	N0.	IÓ	<u>t 8,</u>	8.8	8 8	101	102	105	601	112	113	411	911 911	211	811	120	123	2,3887	2389	2300	2402	2403	2400	7 Cultures 2

GENETICS 3: 169 Mr 1918

TABLE II (continued)

(1) $\frac{w^e r_b f}{c_t}$ (continued)

Coincidence	Coincidence of $w^e r_b$ and fc_t			0	1.7500	0	0 0 0000	2.0000	0 0	и т 1874	0	0	0	0	0	0	0	0/0	0	0	0	ິ	2.8750	0	0	0/0	0	2.6000	
	Total males		46	8	112	52	53	с У	84	6	5 0	25	22	101	91	20	62	50	<u>6</u>	67	62	6 ,	40	47	41	41	64	52	8721
	23	r,			_																								10
		w ^e fcr																							•				~
	23	fc_t				н			۰		•		I	0	7	Ι	0		0	0	I		0	7					82
		20° rb	н ,	-	10	н		(N		~	н	н	T		н	I		H	I	1	I				10	н		911
	8	$r_{\rm b}fc_1$			I																I							I	13
		w					,	-		-	(,	14
		$r_{ m b}c_{ m f}$,	-		. 0		I	I	I								I		I				45
les	н	$v^{e}f$			I				•	-	I	1	I		I		I				I	-					I		8
Ma	ε	۲ ۲	с С	0	ŝ	N	00	، دە	4 (9 0	i u	ο un	ŝ	œ	ŝ	4	ę	61	N	I	Ŋ	ы		ŝ	I	Ι		б	284
		$w^{e}r_{b}fc$		0	ы	I	,	- •	-	л r	. ~	о <i>с</i> о	,	I	ŝ	I	I	I	I	61	H	H	I	н	I		н	н	183
	2	f	18	23	8	15	21	1	5 12	5 Z	ŕœ	21	18	24	22	20	6	12	17	12	61	9	11	11	ю	13	14	11	1854
		20. er b C t	14	II	61	9	12		<u>.</u>	<u>,</u> v	, 5	90	15	14	13	12	6	12	H	×	9	× ×	9	10	6	v	10	9	1440
		$r_{\rm b}f$		ы	4	0	ς, γ	- •		- ~	\$	I	I	I	0	0	I		н			I	I				I	I	163
	ц	$w^e c_f$	10	I	61		ы (N	¢	4 6	7	+	н		I		н		H	-	Ħ		I	N			61	0	153
	1	ct.	31	23	<i>х</i> г	×	10	o i	1	4 v	01	18	19	21	15	13	15	6	12	21	13	×	14	7	7	10	14	9	1884
		$w^{e}r_{b}f$	28	21	30	16	II II	9	4. i	<u>n</u> or	23	? ର୍ଷ	16	<u>5</u> 8	80	25	61	14	12	61	13	12	x	II	17	10	20	21	2406
	Females		137	93	156	59	03	64	020	8 5	100		·8	110	106	88	92	50	æ	75	57	<u>6</u> ,	63	64	<u>6</u> ,	65	8	61	10330
	Culture No.	.041	2503	2504	2505	500	508	015	513	532 232	534	535	536	537	538	539	540	542	543	545	547	503	511	512	540	548	562	292	Totals

GENETICS 3: 170 Mr 1918

TABLE II (continued)

4	C1
e j	r,
A2	
(8	ì

Coincidence	of for	w rb allu Jer	0	0/0	0	1.4667	0	0	0	
	Total males		53	47	20	132	40	52	50	433
	I 2 3	$w^{e}r_{b}fc_{r}$ +				I			I	7
	2 3	$w^{e} r_{b} f c_{r}$	Ι		I	3	II		2	6 7
	ŝ	fcr								
	ц	20'e r _b				61				ġ
les	I 2	$w^{\rm e}r_{\rm b}f$ cr				H	61	-	I	4
Ma	3	v ^b	e	I		ŝ	4	6	I	14
		w ^e fc _r	ъ	0	ę	ъ	н			16
	2	$r_{\rm b}f$	~	12	16	8	×	12	10	94
		$w^{e}c_{f}$	6	<u>м</u>	0	23	ŝ	13	5	8
		f	ę		I	ŝ	н	ы		IO
	I	WerbCr	0		61	ŝ			0	6
		rbcr	10	IO	01	27	Ŋ	9	9	74
		$w^{\rm e}f$	13	17	24	%	14	17	22	135
-	Females		84	61	129	132	53	65	71	595
	Culture No.	418	419	422	440	451	455	406	Totals	

GENETICS 3: 171 Mr 1918

A. Weinstein, Coincidence of crossing over in Drosophila melanogaster (ampelophila)

TABLE II (continued)

$w^e r$,	fr.	1.1
$\left(\cdot \right)$	(3)	

Coincidence	of ant w and fo	a la ma la la	0	1.0007	0	0	0	1.5000	0	4.5000	0	0	0.8750	0	0	2.4167	0	0	0	0	
-	l otal males		12	25	×3	63	51	5	54	45	43	4	40	43	4 <u>5</u>	5 23	90	64	52	19	6101
	1 2 3	w ^e cr r _b f			Ι						-		I								4
	2 3	$w^{\mathbf{e}}r_{\mathbf{h}}f$ $c_{\mathbf{r}}$	ę		2 I	I	I	0	N			I I	°1	I I	I	I	F	I I	I I	I	12 14
	3	rucr		I						-											0
		$w^{e}f$						I		0			1			I					ν.
		$r_{\rm b} f c_{\rm r}$			I				I						I			I			4
les	1 2	we		С	ŝ	I		I	I	I	0	I	I		I		1			0	18
Ma	<u> </u>	f	3	ς		ഗ	I	ন	01	I	-	N	4	0	ŝ	 च	0	4	ŝ	0	40
		20 ^e r _b c r	ŝ	I	ы			cı	ę	ы				I	I	I	I	0	0	I	54
		+	13	16	17	12	12	16	01	×	13	IO	15	11	<u>18</u>	11	21	12	18	23	256
		$w^{e}r_{b}fc$	II	10	12	×	CI	6	7	10	7	×	ŝ		Ŋ	4	~~~	12	4	4	133
	1	$r_{\rm b}$	0	ŝ		4	0	0	I			н	4	-		I	I	Ι	I	9	30
	, I	w ^e fc _r	0	ы	ę	0				I							I	щ	I		13
		fcr	14	12	14	18 8	S	9	ŝ	IO	ΙI	10	I	0		11	12	ΓŢ	-œ	ю	158
		w ^e r _b	18	24	22	12	17	11	22	6	6	~	17	5	1.5	ŝ	10	۲. ۲	13	17	300
	Females		73	109	122	87	12	8	26	73	64	50	10	03	8	65	72	.8	89	120	1471
	Culture No.		579	580	581	585	589	594	604	586	587	010	611	613	615	617	618	620	000	803	Totals

GENETICS 3: 172 Mr 1918

TABLE II (continued) (4) $w^e r_b f c_t$

		Males															Coinciden		
Culture No.	Females		1	I		- 	2		3	<u>I 2</u>		<u> </u>	3	2	3	<u>1</u>	2 3	Total males	of
		w ^e r _b fc _f	+	w	$r_{\rm b} f c_{\rm f}$	$w^{\rm e}r_{\rm b}$	fc,	$w^{e}r_{b}f$	C ₁	wefcı	r _b	w ^e Ct	$r_{\rm b}f$	w ^e r _b c ₁	f	w°f	$r_{\rm b} c_t$		w 7b and
			5								1997				. 5		23		
679	77	13	11	I		II	10	I							2		Ξ,	49	0
680	44	4	18			14	11		4						2			53	0/0
681	100	10	28	3	2	25	15	5	_		r				2			91), 0
1683	84	5	20	2		II	11	2	I			1			2			54	0
084	110	11	31	5		22	22	1	1	1	÷			1	2		1. ÷	97	0
080	19	10	19	2	3	19		3	2	· ·	т			1	2	· .		108 2	0
600	107	24	30	5	1	20	13	3	2	2	1	1		- 4 t	Ť	1		100	0
680	59	14	14	Å	7	13	τλ	2	2	2	2	1		1 T	1			83	ň
600	95	22	27	1 4 T	2	20	14	3	2	-	-			2	2		:1	9 03	0
601	: 00	12	21	2	J T	21	TŦ	5	7				т	1. 5	T		· ·	86	1.2286
605	101	TÒ	34	2	Ē.,	10	8	I	2	I	••		~	10	Ì		۰.	- ET 78	0
742	159	10	33	2	t	35	18	5	6	I	I	2	I		I	ļ .	5	125	3.1250
750	127	23	20	1	4	15	13	4	3	I	I		I	2	3	1		90	10.9890
751	124	ğ	27	1.1	•	26	17	4	3				I	-	1		;	87	10.8750
752	125	17	22	2		14	12	8	2		Í			I.	2	· ·		81	0
753	125	20	32	5	3	28	23	2	6		2		I	2	Ĺ		44 1	125	0.9542
Totals	1779	246	421	36	19	325	225	52	44	9	9	2	5	14	23	j	2	1430	

GENETICS 3: Mr 1918

TABLE 11 (continued)

$(5) \quad \frac{w^{e} I_{z}^{*}}{r_{h} f c_{f}}$

								,							Mal	ês es	-						-						
Culture No.	Females			<u> </u>		2			3		4	<u>I 2</u>		I	3	<u>I</u>	4	2	3	2	4		3 4	<u>I 2 3</u>	<u>I 2 4</u>	<u>1 3 4</u>	234	Total males	Coincidence of
	·	w ^e l _z	r _b fc _f	w°r _b fc _t	, l _z	w°fc _t	$r_{\rm b}l_{\rm z}$	w ^e l _z fc	t r _b	w ^e l _z c _t	r _b f	w [°] r _b l _z	fct	w ^e r _b	l,fc	w ^e r _b f	lzci	w ^e 1	r _b lzfc1	w ^e f	$r_{\rm b}l_{\rm z}c_{\rm f}$	w ^e l _z f	rbCr	+	w ^e r _b l _z c ₁	l _z f	$r_{\rm b}l_{\rm z}f$		$w^{e}r_{b}$ and fc_{t}
702 703 704 705 706 743 745 797 812 813 835 836 845 836 845 890 892 901 913 914 919 921 922 923 925 927 928 943 947	$\begin{array}{c} 142\\ 154\\ 15\\ 95\\ 161\\ 97\\ 82\\ 139\\ 194\\ 148\\ 116\\ 128\\ 80\\ 85\\ 67\\ 125\\ 148\\ 73\\ 142\\ 117\\ 143\\ 173\\ 176\\ 137\\ 151\\ 82\\ 139\\ \end{array}$	37 35 20 13 30 31 22 25 26 28 22 18 13 32 21 32 29 40 36 30 19 26 14 37	$\begin{array}{c} 17\\ 24\\ 12\\ 14\\ 29\\ 6\\ 17\\ 11\\ 19\\ 26\\ 11\\ 15\\ 17\\ 10\\ 11\\ 19\\ 22\\ 12\\ 16\\ 17\\ 18\\ 35\\ 23\\ 27\\ 13\\ 17\\ 22\\ \end{array}$	4 1 2 2 2 4 1 1 1 1 1 4 1 2 2 4 1 2 2 1 2 2	3 2 3 3 1 3 3 4 1 3 4 1 1 4 1 1 4 2 1 3 3 3	7 7 4 3 9 4 4 7 6 8 5 8 3 3 1 1 5 5 10 1 5 5 3 4 9 6 3 8	9 5 9 5 8 8 9 9 5 8 8 9 9 13 2 3 6 5 2 2 2 12 6 6 6 5 5 10 3 8 4 8	16 9 14 8 13 8 5 12 13 2 9 6 6 11 2 9 16 10 5 4 7 11	$\begin{array}{c} 16 \\ 11 \\ 16 \\ 12 \\ 20 \\ 13 \\ 10 \\ 12 \\ 24 \\ 12 \\ 12 \\ 12 \\ 12 \\ 12 \\ 12 \\ 12 \\ 13 \\ 10 \\ 19 \\ 21 \\ 10 \\ 10 \\ 14 \\ 15 \\ 17 \\ 23 \\ 14 \\ 12 \\ 5 \\ 16 \end{array}$	2 I I I 4 I 4 I 4 I 3 2 3 3 4 2 5 4 I 3 2 I I I 1 5 4 I 1 3 2 I I I I I I 1 4 I I 1 4 I I 1 4 I I 1 3 2 3 2 I I 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	71422 377266442465433857462	I I I I	I	I I I I I I I I I I I I I I I I I I I	3 4 ř	2 1		I 2 4 3 I 3 I I 1 2 2 2 I 3 2 2 2 2	I 2 I I I 3	I I I I I I I I I I I I	I I I I I I I I I I	I I I I I I I I I I I I I I I 3	I 3 I I I I 3 I I	I	Γ	I	I	120 105 91 67 121 76 77 96 143 96 69 93 77 65 55 113 151 65 87 90 108 128 116 100 84 70 109	0 3.5000 0 0 0 0 0 0 0 0 0 0 0 0
Totals	3409	737	480	43	53	140	181	219	380	50	114	2	I	23	8	4		40	13	22	12	26	18	I	I	3	I	2572	

* $l_{z} = 1$ ozenge.

GENETICS 3: Mr 1918